# Clustering and EM 

CDS, NYU

April 25, 2022

## Logistics

Final exam

- Period: May 15 4:00-5:50pm EST
- Format: in person, closed book
- Coverage: mainly about material from week 6 onwards but can overlap with basic concepts before midterm


## K-means Clustering

## Unsupervised learning

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Formulation Density estimation: $p(x ; \theta)$ (often with latent variables).
Examples - Discover clusters: cluster data into groups.

- Discover factors: project high-dimensional data to a small number of "meaningful" dimensions, i.e. dimensionality reduction.
- Discover graph structures: learn joint distribution of correlated variables, i.e. graphical models.


## Example: Old Faithful Geyser



- Looks like two clusters.
- How to find these clusters algorithmically?


## k-Means: By Example

- Standardize the data.
- Choose two cluster centers.



## $k$-means: by example

- Assign each point to closest center.



## $k$-means: by example

- Compute new cluster centers.



## $k$-means: by example

- Assign points to closest center.



## $k$-means: by example

- Compute cluster centers.



## $k$-means: by example

- Iterate until convergence.



## Suboptimal Local Minimum

- The clustering for $k=3$ below is a local minimum, but suboptimal:


Would be better to have one cluster here

... and two clusters here

## Formalize $k$-Means

- Dataset $\mathcal{D}=\left\{x_{1}, \ldots, x_{n}\right\} \subset X$ where $X=R^{d}$.


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- The centroid of $C_{i}$ is defined to be

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- The $k$-means objective is to minimize the distance between each example and its cluster centroid:

$$
\begin{equation*}
J(c, \mu)=\sum_{i=1}^{n}\left\|x_{i}-\mu_{c_{i}}\right\|^{2} \tag{2}
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## Avoid bad local minima

$k$-means converges to a local minimum.

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Avoid getting stuck with bad local minima:

- Re-run with random initial centroids.


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- Sequentially choose subsequent centroids from points that are farther away from current centroids:
- Compute distance between each $x_{i}$ and the closest already chosen centroids.
- Randomly choose next centroid with probability proportional to the computed distance squared.


## Summary

We've seen

- Clustering—an unsupervised learning problem that aims to discover group assignments.
- $k$-means:
- Algorithm: alternating between assigning points to clusters and computing cluster centroids.
- Objective: minmizing some loss function by cooridinate descent.
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Next, probabilistic model of clustering.

- A generative model of $x$.
- Maximum likelihood estimation.


## Gaussian Mixture Models

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## Example:

(1) Choose $z \in\{1,2,3\}$ with $p(1)=p(2)=p(3)=\frac{1}{3}$.
(2) Choose $x \mid z \sim \mathcal{N}\left(X \mid \mu_{z}, \Sigma_{z}\right)$.


## Gaussian mixture model (GMM)

Generative story of GMM with $k$ mixture components:
(1) Choose cluster $z \sim \operatorname{Categorical}\left(\pi_{1}, \ldots, \pi_{k}\right)$.
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Probability density of $x$ :

- Sum over (marginalize) the latent variable $z$.

$$
\begin{align*}
p(x) & =\sum_{z} p(x, z)  \tag{5}\\
& =\sum_{z} p(x \mid z) p(z)  \tag{6}\\
& =\sum_{k} \pi_{k} \mathcal{N}\left(\mu_{k}, \Sigma_{k}\right) \tag{7}
\end{align*}
$$

## Identifiability Issues for GMM

- Suppose we have found parameters

$$
\begin{aligned}
\text { Cluster probabilities: } & \pi=\left(\pi_{1}, \ldots, \pi_{k}\right) \\
\text { Cluster means: } & \mu=\left(\mu_{1}, \ldots, \mu_{k}\right) \\
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- Assuming all clusters are distinct, there are $k$ ! equivalent solutions.
- Not a problem per se, but something to be aware of.


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- MLE (also called maximize marginal likelihood).
- Log likelihood of data:

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L(\theta) & =\sum_{i=1}^{n} \log p\left(x_{i} ; \theta\right)  \tag{8}\\
& =\sum_{i=1}^{n} \log \sum_{z} p(x, z ; \theta) \tag{9}
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- Cannot push $\log$ into the sum... $z$ and $x$ are coupled.
- No closed-form solution for GMM—try to compute the gradient yourself!


## Gradient Descent / SGD for GMM

- What about running gradient descent or SGD on

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J(\pi, \mu, \Sigma)=-\sum_{i=1}^{n} \log \left\{\sum_{z=1}^{k} \pi_{z} \mathcal{N}\left(x_{i} \mid \mu_{z}, \Sigma_{z}\right)\right\} ?
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- Then $\Sigma_{i}$ is positive semidefinite.
- Even then, pure gradient-based methods have trouble. ${ }^{1}$
${ }^{1}$ See Hosseini and Sra's Manifold Optimization for Gaussian Mixture Models for discussion and further references.


## Learning GMMs: observable case

Suppose we observe cluster assignments $z$. Then MLE is easy:

$$
\begin{align*}
n_{z} & =\sum_{i=1}^{n} 1\left(z_{i}=z\right) & & \text { \# examples in each cluster }  \tag{10}\\
\hat{\pi}(z) & =\frac{n_{z}}{n} & & \text { fraction of examples in each cluster }  \tag{11}\\
\hat{\mu}_{z} & =\frac{1}{n_{z}} \sum_{i: z_{i}=z} x_{i} & & \text { empirical cluster mean }  \tag{12}\\
\hat{\Sigma}_{z} & =\frac{1}{n_{z}} \sum_{i: z_{i}=z}\left(x_{i}-\hat{\mu}_{z}\right)\left(x_{i}-\hat{\mu}_{z}\right)^{T} . & & \text { empirical cluster covariance } \tag{13}
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& =\frac{\pi_{j} \mathcal{N}\left(x_{i} \mid \mu_{j}, \Sigma_{j}\right)}{\sum_{k} \pi_{k} \mathcal{N}\left(x_{i} \mid \mu_{k}, \Sigma_{k}\right)} \tag{16}
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- $p(z \mid x)$ is a soft assignment.
- If we know the parameters $\mu, \Sigma, \pi$, this would be easy to compute.


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- compute soft assignments $p\left(z \mid x_{i}\right)$ for all $i$.
(2) M-step: standard MLE for $\mu, \Sigma, \pi$ given "observed" variables.
- Equivalent to MLE in the observable case on data weighted by $p\left(z \mid x_{i}\right)$.


## M-step for GMM

- Let $p(z \mid x)$ be the soft assignments:

$$
\gamma_{i}^{j}=\frac{\pi_{j}^{\text {old }} \mathcal{N}\left(x_{i} \mid \mu_{j}^{\text {old }}, \Sigma_{j}^{\text {old }}\right)}{\sum_{c=1}^{k} \pi_{c}^{\text {old }} \mathcal{N}\left(x_{i} \mid \mu_{c}^{\text {old }}, \Sigma_{c}^{\text {old }}\right)} .
$$

- Exercise: show that

$$
\begin{aligned}
n_{z} & =\sum_{i=1}^{n} \gamma_{i}^{z} \\
\mu_{z}^{\text {new }} & =\frac{1}{n_{z}} \sum_{i=1}^{n} \gamma_{i}^{z} x_{i} \\
\Sigma_{z}^{\text {new }} & =\frac{1}{n_{z}} \sum_{i=1}^{n} \gamma_{i}^{z}\left(x_{i}-\mu_{z}^{\text {new }}\right)\left(x_{i}-\mu_{z}^{\text {new }}\right)^{T} \\
\pi_{z}^{\text {new }} & =\frac{n_{z}}{n} .
\end{aligned}
$$

## EM for GMM

- Initialization



## EM for GMM

- First soft assignment:



## EM for GMM

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## EM for GMM

- After 5 rounds of EM:



## EM for GMM

- After 20 rounds of EM:



## EM for GMM: Summary

- EM is a general algorithm for learning latent variable models.
- Key idea: if data was fully observed, then MLE is easy.
- E-step: fill in latent variables by computing $p(z \mid x, \theta)$.
- M-step: standard MLE given fully observed data.
- Simpler and more efficient than gradient methods.
- Can prove that EM monotonically improves the likelihood and converges to a local minimum.
- $k$-means is a special case of EM for GMM with hard assignments, also called hard-EM.


## Latent Variable Models

## General Latent Variable Model

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## Definition

A latent variable model is a probability model for which certain variables are never observed.

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A latent variable model is a probability model for which certain variables are never observed.
e.g. The Gaussian mixture model is a latent variable model.

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- To simplify notation, take $x$ to represent the entire dataset

$$
x=\left(x_{1}, \ldots, x_{n}\right),
$$

and $z$ to represent the corresponding unobserved variables

$$
z=\left(z_{1}, \ldots, z_{n}\right)
$$

- An observation of $x$ is called an incomplete data set.
- An observation $(x, z)$ is called a complete data set.


## Our Objectives

- Learning problem: Given incomplete dataset $x$, find MLE

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- For Gaussian mixture model, learning is hard, inference is easy.
- For more complicated models, inference can also be hard. (See DSGA-1005)

Log-Likelihood and Terminology

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- We often call $p(x, z)$ the joint. (for "joint distribution")
- Similarly, $\log p(x)$ is the marginal log-likelihood.


## EM Algorithm

## Intuition

Problem: marginal $\log$-likelihood $\log p(x ; \theta)$ is hard to optimize (observing only $x$ )
Observation: complete data $\log$-likelihood $\log p(x, z ; \theta)$ is easy to optimize (observing both $x$ and $z$ )

Idea: guess a distribution of the latent variables $q(z)$ (soft assignments)
Maximize the expected complete data log-likelihood:

$$
\max _{\theta} \sum_{z \in z} q(z) \log p(x, z ; \theta)
$$

EM assumption: the expected complete data log-likelihood is easy to optimize Why should this work?

## Math Prerequisites

## Jensen's Inequality

## Theorem (Jensen's Inequality)

If $f: \mathrm{R} \rightarrow \mathrm{R}$ is a convex function, and $x$ is a random variable, then

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Moreover, if $f$ is strictly convex, then equality implies that $x=\mathbb{E} x$ with probability 1 (i.e. $x$ is a constant).

- e.g. $f(x)=x^{2}$ is convex. So $\mathbb{E} x^{2} \geqslant(\mathbb{E} x)^{2}$. Thus

$$
\operatorname{Var}(x)=\mathbb{E} x^{2}-(\mathbb{E} x)^{2} \geqslant 0
$$

## Kullback-Leibler Divergence

- Let $p(x)$ and $q(x)$ be probability mass functions (PMFs) on $X$.
- How can we measure how "different" $p$ and $q$ are?


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- The Kullback-Leibler or "KL" Divergence is defined by

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\operatorname{KL}(p \| q)=\sum_{x \in X} p(x) \log \frac{p(x)}{q(x)}
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(Assumes $q(x)=0$ implies $p(x)=0$.)

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- Can also write this as

$$
\operatorname{KL}(p \| q)=\mathbb{E}_{x \sim p} \log \frac{p(x)}{q(x)}
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## Gibbs Inequality $(\operatorname{KL}(p \| q) \geqslant 0$ and $\operatorname{KL}(p \| p)=0)$

## Theorem (Gibbs Inequality)

Let $p(x)$ and $q(x)$ be PMFs on $X$. Then

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K L(p \| q) \geqslant 0
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with equality iff $p(x)=q(x)$ for all $x \in \mathcal{X}$.

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- KL divergence measures the "distance" between distributions.
- Note:
- KL divergence not a metric.
- KL divergence is not symmetric.


## Gibbs Inequality: Proof

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- Since $-\log$ is strictly convex, we have strict equality iff $q(x) / p(x)$ is a constant, which implies $q=p$.

The ELBO: Family of Lower Bounds on $\log p(x \mid \theta)$

The Maximum Likelihood Estimator


Lower bound of the marginal log-likelihood

$$
\log p(x ; \theta)=\log \sum_{z \in Z} p(x, z ; \theta)
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& \geqslant \sum_{z \in z} q(z) \log \frac{p(x, z ; \theta)}{q(z)} \\
& \stackrel{\text { def }}{=} \mathcal{L}(q, \theta)
\end{aligned}
$$

- Evidence: $\log p(x ; \theta)$
- Evidence lower bound (ELBO): $\mathcal{L}(q, \theta)$
- $q$ : chosen to be a family of tractable distributions
- Idea: maximize the $E L B O$ instead of $\log p(x ; \theta)$


## MLE, EM, and the ELBO

- The MLE is defined as a maximum over $\theta$ :

$$
\hat{\theta}_{\text {MLE }}=\underset{\theta}{\arg \max }[\log p(x \mid \theta)] .
$$

- For any PMF $q(z)$, we have a lower bound on the marginal log-likelihood

$$
\log p(x \mid \theta) \geqslant \mathcal{L}(q, \theta) .
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- In EM algorithm, we maximize the lower bound (ELBO) over $\theta$ and $q$ :

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- In EM algorithm, q ranges over all distributions on $z$.


## EM: Coordinate Ascent on Lower Bound

- Choose sequence of $q$ 's and $\theta$ 's by "coordinate ascent" on $\mathcal{L}(q, \theta)$.


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(3) Let $\theta^{\text {new }}=\arg \max _{\theta} \mathcal{L}\left(q^{*}, \theta\right)$.
(9) Go to step 2 , until converged.
- Will show: $p\left(x \mid \theta^{\text {new }}\right) \geqslant p\left(x \mid \theta^{\text {old }}\right)$
- Get sequence of $\theta$ 's with monotonically increasing likelihood.


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Is ELBO a "good" lowerbound?

$$
\begin{aligned}
\mathcal{L}(q, \theta) & =\sum_{z \in \mathcal{Z}} q(z) \log \frac{p(x, z \mid \theta)}{q(z)} \\
& =\sum_{z \in \mathcal{Z}} q(z) \log \frac{p(z \mid x, \theta) p(x \mid \theta)}{q(z)} \\
& =-\sum_{z \in \mathcal{Z}} q(z) \log \frac{q(z)}{p(z \mid x, \theta)}+\sum_{z \in \mathcal{Z}} q(z) \log p(x \mid \theta) \\
& =-\operatorname{KL}(q(z) \| p(z \mid x, \theta))+\underbrace{\log p(x \mid \theta)}_{\text {evidence }}
\end{aligned}
$$

- KL divergence: measures "distance" between two distributions (not symmetric!)
- $\operatorname{KL}(q \| p) \geqslant 0$ with equality iff $q(z)=p(z \mid x)$.
- $\mathrm{ELBO}=$ evidence $-\mathrm{KL} \leqslant$ evidence


## Maximizing over $q$ for fixed $\theta$.

- Find q maximizing

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\mathcal{L}(q, \theta)=-\operatorname{KL}[q(z), p(z \mid x, \theta)]+\log p(x \mid \theta)
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- Summary:

$$
\log p(x \mid \theta)=\sup _{q} \mathcal{L}(q, \theta) \quad \forall \theta
$$

- For any $\theta$, sup is attained at $q(z)=p(z \mid x, \theta)$.

Marginal Log-Likelihood IS the Supremum over Lower Bounds


## Summary

Latent variable models: clustering, latent structure, missing lables etc.
Parameter estimation: maximum marginal log-likelihood
Challenge: directly maximize the evidence $\log p(x ; \theta)$ is hard
Solution: maximize the evidence lower bound:

$$
\mathrm{ELBO}=\mathcal{L}(q, \theta)=-\mathrm{KL}(q(z) \| p(z \mid x ; \theta))+\log p(x ; \theta)
$$

Why does it work?

$$
\begin{aligned}
q^{*}(z) & =p(z \mid x ; \theta) \quad \forall \theta \in \Theta \\
\mathcal{L}\left(q^{*}, \theta^{*}\right) & =\max _{\theta} \log p(x ; \theta)
\end{aligned}
$$

## EM algorithm

Coordinate ascent on $\mathcal{L}(q, \theta)$
(1) Random initialization: $\theta^{\text {old }} \leftarrow \theta_{0}$
(2) Repeat until convergence
(1) $q(z) \leftarrow \arg \max _{q} \mathcal{L}\left(q, \theta^{\text {old }}\right)$

$$
\text { Expectation (the E-step): } \quad \begin{aligned}
q^{*}(z) & =p\left(z \mid x ; \theta^{\text {old }}\right) \\
J(\theta) & =\mathcal{L}\left(q^{*}, \theta\right)
\end{aligned}
$$

(1) $\theta^{\text {new }} \leftarrow \arg \max _{\theta} \mathcal{L}\left(q^{*}, \theta\right)$

Maximization (the M-step): $\quad \theta^{\text {new }} \leftarrow \underset{\theta}{\arg \max } J(\theta)$

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[Equivalent to maximizing expected complete log-likelihood.]

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## (2) Maximization Step

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[Equivalent to maximizing expected complete log-likelihood.]
EM puts no constraint on $q$ in the E-step and assumes the M -step is easy. In general, both steps can be hard.

## Monotonically increasing likelihood



Exercise: prove that EM increases the marginal likelihood monotonically $^{\theta^{\text {ood }}}$

$$
\log p\left(x ; \theta^{\text {new }}\right) \geqslant \log p\left(x ; \theta^{\text {old }}\right)
$$

Does EM converge to a global maximum?

## Variations on EM

## EM Gives Us Two New Problems

- The "E" Step: Computing

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J(\theta):=\mathcal{L}\left(q^{*}, \theta\right)=\sum_{z} q^{*}(z) \log \left(\frac{p(x, z \mid \theta)}{q^{*}(z)}\right)
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- The "M" Step: Computing

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\theta^{\text {new }}=\underset{\theta}{\arg \max } J(\theta) .
$$

- Either of these can be too hard to do in practice.


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- e.g. take a gradient step on J.
- We still get monotonically increasing likelihood.


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- Solution: Restrict to distributions $Q$ that are easy to work with.
- Lower bound now looser:

$$
q^{*}=\underset{q \in \mathcal{Q}}{\arg \min } \operatorname{KL}\left[q(z), p\left(z \mid x, \theta^{\text {old }}\right)\right]
$$

## Today's Summary

- Motivation: Unsupervised learning
- K-means: A simple algorithm for discovering clusters
- Making k-means probabilistic: Gaussian mixture models
- More generally: Latent variable models
- Learning of latent variable models: EM
- Underlying principle: Maximizing ELBO


[^0]:    ${ }^{1}$ See Hosseini and Sra's Manifold Optimization for Gaussian Mixture Models for discussion and further references.

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