

# Recitation 8

## Bayesian Methods

DS-GA 1003 Machine Learning

CDS

March 21, 2023

# Agenda

- 1 Recap: MLE
- 2 Bayesian Methods
- 3 Questions

# MLE for Conditional Probability Models

- Observed data  $\mathcal{D} = \{x_{1\dots n}, y_{1\dots n}\}$
- Compute likelihood of the data as a function of parameter(s)  $\theta$

$$L_{\mathcal{D}}(\theta) = \prod_{i=1}^n p(y_i|x_i; \theta)$$

- Find that value of  $\theta \in \Theta$  which maximizes the likelihood  $\rightarrow$  MLE
  - MLE is the ERM of NLL loss

$$\hat{\theta}_{MLE} = \arg \max_{\theta} \prod_{i=1}^n p(y_i|x_i; \theta)$$

- And we make predictions on new points  $x'$  as:

$$\hat{f}(x') = p(y|x'; \hat{\theta}_{MLE})$$

# MLE for Conditional Probability Models

- Observe that  $\hat{\theta}_{MLE}$  is very dependent on the observed data
- Can we do better? What if you have an intuition/belief about the parameter  $\theta$  before observing the data  $\mathcal{D}$ ?

# Bayesian Methods

- Ingredients:
  - **Parameter space**  $\Theta$ .
  - **Prior**: Distribution  $p(\theta)$  on  $\Theta$ .
  - **Action space**  $\mathcal{A}$ .
  - **Loss function**:  $\ell : \mathcal{A} \times \Theta \rightarrow \mathbb{R}$ .

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  - **Loss function**:  $\ell : \mathcal{A} \times \Theta \rightarrow \mathbb{R}$ .
- The prior  $p(\theta)$  represents your belief about the parameter without seeing the data
- And you update this belief based on observing the data  $\mathcal{D}$  with Bayes rule to get the posterior
- Posterior  $p(\theta|D) \propto p(\mathcal{D}|\theta)p(\theta)$
- From this distribution, we can get point estimates or take actions

# Bayesian Decision Theory

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  - **Parameter space**  $\Theta$ .
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  - **Action space**  $\mathcal{A}$ .
  - **Loss function**:  $\ell : \mathcal{A} \times \Theta \rightarrow \mathbb{R}$ .
- The **posterior risk** of an action  $a \in \mathcal{A}$  is

$$\begin{aligned}r(a) &:= \mathbb{E}[\ell(\theta, a) \mid \mathcal{D}] \\ &= \int \ell(\theta, a)p(\theta \mid \mathcal{D}) d\theta.\end{aligned}$$

- It's the **expected loss under the posterior**.



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- It's the **expected loss under the posterior**.
- A **Bayes action**  $a^*$  is an action that minimizes posterior risk:

$$r(a^*) = \min_{a \in \mathcal{A}} r(a)$$

# MAP Estimator

- How do we predict  $y$  from the posterior of  $\theta$ ?
- MAP estimator for  $\theta$  from the posterior

$$\hat{\theta}_{MAP} = \arg \max_{\theta} p(\theta | \mathcal{D})$$

- We can predict  $y$  by

$$\hat{y} = \arg \max_y p(y | x; \theta = \hat{\theta}_{MAP})$$

# The Posterior Predictive Distribution

- The MAP estimator only depends on the **mode** of the posterior. Is there a way to incorporate all the information from the posterior?
- The **posterior predictive distribution** is given by

$$p(y | x, \mathcal{D}) = \int p(y | x; \theta) p(\theta | \mathcal{D}) d\theta.$$

- This is an average of all conditional densities in our family, weighted by the posterior.
- May not have closed form. Numerical integral may be hard to compute.

# MAP Estimator vs Posterior Predictive Distribution

- How do we predict by posterior predictive distribution given a new data point?

$$\hat{y} = \arg \max_y p(y | x, \mathcal{D}) = \arg \max_y \int p(y | x; \theta) p(\theta | \mathcal{D}) d\theta.$$

- Different to the MAP estimator:

$$\hat{\theta}_{MAP} = \arg \max_{\theta} p(\theta | \mathcal{D})$$

$$\hat{y} = \arg \max_y p(y | x; \theta = \hat{\theta}_{MAP})$$

- In general, the predictions from two methods are different.

# MAP Estimator Vs MLE

- MLE looks for the value that maximizes likelihood alone

$$\hat{\theta}_{MLE} = \arg \max_{\theta} L_{\mathcal{D}}(\theta) = \arg \max_{\theta} \prod_{i=1}^n p(y_i | x_i; \theta)$$

- MAP maximizes the posterior i.e. a combination of prior and likelihood

$$\hat{\theta}_{MAP} = \arg \max_{\theta} p(\theta | \mathcal{D}) = \arg \max_{\theta} L_{\mathcal{D}}(\theta)p(\theta)$$

## Question 1

**Question 1.** (From DeGroot and Schervish) Let  $\theta$  denote the proportion of registered voters in a large city who are in favor of a certain proposition. Suppose that the value of  $\theta$  is unknown, and two statisticians  $A$  and  $B$  assign to  $\theta$  the following different (beta) prior PDFs  $\xi_A(\theta)$  and  $\xi_B(\theta)$ , respectively:

$$\begin{aligned}\xi_A(\theta) &= 2\theta && \text{for } 0 < \theta < 1, \\ \xi_B(\theta) &= 4\theta^3 && \text{for } 0 < \theta < 1.\end{aligned}$$

In a random sample of 1000 registered voters from the city, it is found that 710 are in favor of the proposition.

- Find the posterior distribution that each statistician assigns to  $\theta$ .

## Question 1: Solution

- Likelihood of the observed data, 710 in-favour, 290 against:

$$f(x|\theta) = \theta^{710}(1 - \theta)^{290}$$

- Multiplying by the two priors  $\xi_A$  and  $\xi_B$ , we have

$$\xi_A(\theta|x) \propto f(x|\theta)\xi_A(\theta) \propto \theta^{711}(1 - \theta)^{290}$$

and

$$\xi_B(\theta|x) \propto f(x|\theta)\xi_B(\theta) \propto \theta^{713}(1 - \theta)^{290}.$$

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- Thus the posteriors from  $A$  and  $B$  are both beta with parameters  $(712, 291)$  and  $(714, 291)$ , respectively.



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In a random sample of 1000 registered voters from the city, it is found that 710 are in favor of the proposition.

- Find the Bayes estimate of  $\theta$  (minimizer of posterior expected loss) for each statistician based on the squared error loss function.

## Question 1: Solution

If the loss function is square loss, the minimizer  $f^* = E[Y|X]$ .

- We have found the two posteriors  $\xi_A(\theta|x)$  and  $\xi_B(\theta|x)$
- The posteriors from  $A$  and  $B$  are both beta with parameters  $(712, 291)$  and  $(714, 291)$ , respectively.

## Question 1: Solution

If the loss function is square loss, the minimizer  $f^* = E[Y|X]$ .

- We have found the two posteriors  $\xi_A(\theta|x)$  and  $\xi_B(\theta|x)$
- The posteriors from  $A$  and  $B$  are both beta with parameters  $(712, 291)$  and  $(714, 291)$ , respectively.
- Thus minimizers of the posterior expected loss is the respective means are  $\frac{712}{1003}$  and  $\frac{714}{1005}$ .
  - Recall the mean of a Beta distribution  $\mathbb{E}[x; a, b] = \frac{a}{a+b}$

## Question 2

What would be the Maximum a Posteriori (MAP) estimator for  $\lambda$  for i.i.d.  $\{x_1, x_2, \dots, x_N\}$  where  $x_i \sim \exp(\lambda)$  with prior  $\lambda \sim \text{Uniform}[u_0, u_1]$ ?

## Question 2: Solution

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- Likelihood:  $L(x_1, \dots, x_N | \lambda) = \lambda^N e^{-\lambda(x_1 + \dots + x_N)}$
- log-likelihood:  $\ell(\lambda | x_1, \dots, x_N) = N \ln \lambda - \lambda(x_1 + \dots + x_N)$
- $\ell'(\lambda) =$ 

$$\frac{N}{\lambda} - (x_1 + \dots + x_N) \begin{cases} > 0 & \text{if } 0 < \lambda < 1/\bar{x} = N/(x_1 + \dots + x_N), \\ = 0 & \text{if } \lambda = 1/\bar{x} \\ < 0 & \text{if } \lambda > 1/\bar{x} \end{cases}$$
- Prior:  $p(\lambda) = \frac{1}{u_1 - u_0} \mathbb{1}_{[u_0, u_1]}(\lambda)$ .
- Posterior:
 
$$p(\lambda | x_1, \dots, x_N) \propto L(x_1, \dots, x_N | \lambda) p(\lambda) = \lambda e^{-\lambda(x_1 + \dots + x_N)} \mathbb{1}_{[u_0, u_1]}(\lambda)$$
- Maximum value of posterior is attained at

$$\lambda = \begin{cases} u_0 & \text{if } u_0 > 1/\bar{x}, \\ 1/\bar{x} & \text{if } u_0 \leq 1/\bar{x} \leq u_1 \\ u_1 & \text{if } u_1 < 1/\bar{x}. \end{cases}$$

# Takeaways

- In Bayesian methods, we have a prior that encodes our belief without the data
- We update the prior based on the observed data i.e. likelihood and get the posterior distribution
- What can we do with this distribution? MAP estimator, Bayesian point estimation, credible set, etc.